

2021.3 Question 11

1. From the definitions, $X \sim \text{Exp}(\lambda)$, and $Y = \lfloor X \rfloor$.

Hence, for $n \geq 0$,

$$\begin{aligned}
 P(Y = n) &= P(\lfloor X \rfloor = n) \\
 &= P(n \leq X < n + 1) \\
 &= \int_n^{n+1} f(x) \, dx \\
 &= \int_n^{n+1} \lambda \cdot e^{-\lambda x} \, dx \\
 &= [-e^{-\lambda x}]_n^{n+1} \\
 &= -e^{-\lambda(n+1)} + e^{-\lambda n} \\
 &= e^{-n\lambda} (1 - e^{-\lambda}),
 \end{aligned}$$

as desired.

2. Since $Z = X - Y$, we know that $Z = \{X\}$ where $\{x\}$ stands for the fractional part of x .

Hence, for $0 \leq z \leq 1$, we have

$$\begin{aligned}
 P(Z < z) &= P(\{X\} < z) \\
 &= P(X - Y < z) \\
 &= \sum_{n=0}^{\infty} P(X < Y + z, Y = n) \\
 &= \sum_{n=0}^{\infty} P(n \leq X < n + z) \\
 &= \sum_{n=0}^{\infty} \int_n^{n+z} \lambda \cdot e^{-\lambda x} \, dx \\
 &= \sum_{n=0}^{\infty} [-e^{-\lambda x}]_n^{n+z} \\
 &= \sum_{n=0}^{\infty} [-e^{-\lambda(n+z)} + e^{-\lambda n}] \\
 &= \sum_{n=0}^{\infty} e^{-n\lambda} (1 - e^{-\lambda z}) \\
 &= (1 - e^{-\lambda z}) \sum_{n=0}^{\infty} e^{-n\lambda} \\
 &= (1 - e^{-\lambda z}) \cdot \frac{1}{1 - e^{-\lambda}} \\
 &= \frac{1 - e^{-\lambda z}}{1 - e^{-\lambda}},
 \end{aligned}$$

as desired.

3. It must be the case that $0 \leq Z < 1$, and the cumulative distribution function of Z is given by, for $0 \leq z \leq 1$,

$$F_Z(z) = \frac{1 - e^{-\lambda z}}{1 - e^{-\lambda}}.$$

By differentiating with respect to z , we get the probability density function of Z is given by, for

$$0 \leq z \leq 1,$$

$$\begin{aligned} f_Z(z) &= F'_Z(z) \\ &= \frac{d}{dz} \frac{1 - e^{-\lambda z}}{1 - e^{-\lambda}} \\ &= \frac{1}{1 - e^{-\lambda}} \cdot (\lambda \cdot e^{-\lambda z}) \\ &= \frac{\lambda e^{-\lambda z}}{1 - e^{-\lambda}}, \end{aligned}$$

and zero everywhere else.

Hence, the expectation is given by

$$\begin{aligned} E(Z) &= \int_0^1 z f_Z(z) dz \\ &= \int_0^1 \frac{\lambda z e^{-\lambda z}}{1 - e^{-\lambda}} dz \\ &= \frac{\lambda}{1 - e^{-\lambda}} \int_0^1 z e^{-\lambda z} dz \\ &= -\frac{1}{1 - e^{-\lambda}} \int_0^1 z d e^{-\lambda z} \\ &= -\frac{1}{1 - e^{-\lambda}} \left[(z e^{-\lambda z})_0^1 - \int_0^1 e^{-\lambda z} dz \right] \\ &= -\frac{1}{1 - e^{-\lambda}} \left[z e^{-\lambda z} + \frac{e^{-\lambda z}}{\lambda} \right]_0^1 \\ &= -\frac{1}{1 - e^{-\lambda}} \left[\left(e^{-\lambda} + \frac{e^{-\lambda}}{\lambda} \right) - \left(0 + \frac{1}{\lambda} \right) \right] \\ &= \frac{\frac{1}{\lambda} - \frac{e^{-\lambda}}{\lambda} - e^{-\lambda}}{1 - e^{-\lambda}} \\ &= \frac{1 - e^{-\lambda} - \lambda e^{-\lambda}}{\lambda(1 - e^{-\lambda})}. \end{aligned}$$

4. Since $0 \leq z_1 < z_2 \leq 1$, we have $n \leq n + z_1 < n + z_2 \leq n + 1$, and hence

$$\begin{aligned} P(Y = n, z_1 < Z < z_2) &= P(Y = n, z_1 < X - Y < z_2) \\ &= P(n + z_1 < X < n + z_2) \\ &= \int_{n+z_1}^{n+z_2} \lambda \cdot e^{-\lambda x} \\ &= [-e^{-\lambda x}]_{n+z_1}^{n+z_2} \\ &= e^{-\lambda(n+z_1)} - e^{-\lambda(n+z_2)} \\ &= e^{-\lambda n} [e^{-\lambda z_1} - e^{-\lambda z_2}]. \end{aligned}$$

On the other hand, notice

$$\begin{aligned} P(Y = n) P(z_1 < Z < z_2) &= P(Y = n) (P(Z < z_2) - P(Z < z_1)) \\ &= (1 - e^{-\lambda}) e^{-n\lambda} \cdot \left[\frac{1 - e^{-\lambda z_2}}{1 - e^{-\lambda}} - \frac{1 - e^{-\lambda z_1}}{1 - e^{-\lambda}} \right] \\ &= e^{-n\lambda} [(1 - e^{-\lambda z_2}) - (1 - e^{-\lambda z_1})] \\ &= e^{-n\lambda} [e^{-\lambda z_1} - e^{-\lambda z_2}]. \end{aligned}$$

Hence, we have

$$P(Y = n, z_1 < Z < z_2) = P(Y = n) P(z_1 < Z < z_2),$$

and we can conclude that Y and Z are independent.